Journal of Biostatistics and Epidemiology

J Biostat Epidemiol. 2022;8(2):208-233

Original Article

Multi-Class Classification using Mixtures of Univariate and Multivariate ROC Curves

Siva Gajjalavari, Vishnu Vardhan Rudravaram*

Department of Statistics, Pondicherry University, Puducherry, India.

ABSTRACT

ARTICLE INFO

Received 14.02.2022 Revised 07.04.2022 Accepted 06.05.2022 Published 15.07.2022

Keywords:

Mixture models; Bi-Normal ROC; Multivariate ROC; Multi-class classification; Area under the Curve. **Introduction:** Receiver Operating Characteristic (ROC) curve is one of the widely used supervised classification technique to allocate/classify the individuals and also instrumental in comparing diagnostic tests. Generally, to deal with classification problems we need to have knowledge on class labels. In most of the medical scenarios, most of data sets exhibit multi-model patterns in class labels which leads to multi-class classification problems.

The main aim of this study is to address on the issue of constructing ROC models when there exists multimodel patterns in the class labels further, to classify the individuals for better diagnosis and also to reduce the complexity of graphical representation of ROC curves in such classification problems.

Methods: A new version of univariate and multivariate ROC models are proposed in the framework of Finite Mixtures, due to the flexibility of identifying and modelling the subcomponents in the heterogeneous populations.

Results: Oral Glucose Tolerance Test and Disk Hernia datasets are used and simulation studies are also performed. Results show that the proposed models possess better accuracy when compared with Bi-Normal and MROC models with reasonable low 1-Specificity and higher Sensitivity. The ROC curves are depicted in a 2D space rather than higher dimension for multi-class classification problem.

Conclusion: It is suggested that before one proceeds to model ROC curves, it is better to take a look at the density patterns of the study variable(s), which in turn help in explaining the true information between the classes and also provides good amount of "true" accuracy.

Introduction

Over the past seven decades, the Receiver Operating Characteristic (ROC) curve has witnessed developments in both theory and application. Particularly, in the field of diagnostic medicine, ROC curve has been widely used for evaluating the test's

^{*.}Corresponding Author: vrstatsgur@gmail.com



Copyright © 2021 Tehran University of Medical Sciences. Published by Tehran University of Medical Sciences. This work is licensed under a Creative Commons Attribution-NonCommercial 4.0 International license (https://creativecommons.org/licenses/by-nc/4.0/). Noncommercial uses of the work are permitted, provided the original work is properly cited. performance and also useful in comparing diagnostic tests by means of Area under the Curve (AUC) and Sensitivities.¹ ROC curve is a unit square plot between 1-specificity (False Positive Rate) and sensitivity (True Positive Rate) at various thresholds.

On literature review about modelling ROC curve, good number of articles were found, which were based on the assumption that the populations follow Normal and Non-Normal distributions.²⁻⁹ The most widely used parametric ROC model is the Bi-Normal model,² which assumes that both the populations are distributed as Normal. For more details on various bi-distributional ROC models, one can refer to.¹⁰

In the recent past, multivariate extensions of ROC curve were proposed under the assumption that the populations follow multivariate normal.¹¹⁻¹⁵

In general, to deal with classification situations, we need to have knowledge on class labels. Even if the class labels are known, still there might be sub populations within each class labels. For instance, let us consider the OGTT (Oral Glucose Tolerance Test) data,¹⁶ which has 21 sample observations and the histogram of the same is shown in figure 1(a). Here, an attempt is made to see whether there are any hidden sub populations with in healthy and diseased of OGTT data by using the EM algorithm.¹⁷ On such exploration, a bi-modal pattern is witnessed in the diseased population, turning out to a total of three classes in the OGTT data shown in figure 1(b). So, before proceeding to build a classification model, it is better to do such an exercise to explore and find out the hidden sub populations in each of the known class labels, if any. One of the points of exploring the sub populations in each known class label is to further classify the individuals for better diagnosis or treatment regime. Now, the two-class problem extended to three-class problem, in general, it can be referred as a multi-class classification problem.

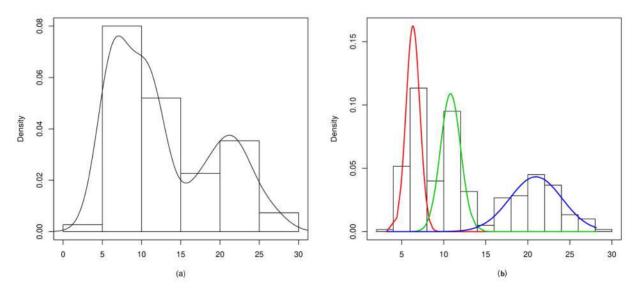


Figure 1. (a) The overall density plot of OGTT, (b) Plot after identifying the components in the OGTT data set. The plots consists of

- (a) Histogram and overlying the density curve of OGTT data.
- (b) Identified sub components in OGTT data and resulted in 3- components.

Generally, the Bi-normal ROC curve and Multivariate Receiver Operating Characteristic Curve (MROC) curve will be useful when we deal with a binary class framework. Hence, there is a need to come out with an ROC curve in both univariate and multivariate setup that can accommodate the multi-class scenario. Here, to have the mathematical flexibility and advantages in constructing ROC curve, the multi-class ROC models are proposed in the framework of Finite Mixture Models (FMM).

Finite Mixture Models

These are widely used in many scientific areas, where the data likely to have several sub populations that are to be determined. FMM provides extreme flexibility in model fitting when the data have many modes, skewness and non-distributional characteristics. For more detailed account of major issues, methodologies on FMM and its applications in diversified areas, readers can see.^{18, 19}

Generally, a p-component mixture density is given by

$$f(x_j; \Psi) = \sum_{i=1}^{p} \lambda_i \phi(x_j; \theta_i)$$

where the vector Ψ contains all unknown parameters of mixture model i.e.,

 $\Psi = (\lambda_1, \lambda_2, \dots, \lambda_{p-1}, \xi')', \ \lambda_i \ 's \text{ are the mixing}$ proportions $(0 < \lambda_i < 1); \sum_{i=1}^p \lambda_i = 1 \text{ and } \xi \text{ is the}$ vector that consists all the distinct parameters in $(\theta_1, \theta_2, \dots, \theta_p)$.

Let us define the density function as $\phi(x_i; \theta_i)$,

$$\phi(x_j; \theta_i) = \begin{cases} \phi(x_j; \mu_i, \sigma_i^2); Normal \\ \phi(x_j; \mu_i, \Sigma_i); Multivariate Normal \end{cases}$$

here,
$$\phi(x_j; \mu_i, \sigma_i^2) = \frac{1}{\sigma_i \sqrt{2\pi}} \exp\left\{-\frac{1}{2} \left(\frac{x_j - \mu_i}{\sigma_i}\right)^2\right\}$$

is the normal density; μ_i and σ_i^2 are mean and variance of ith population.

$$\phi(\mathbf{x}_j; \, \boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i) = \frac{1}{(2\pi)^{p/2} |\boldsymbol{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_i)' \boldsymbol{\Sigma}_i^{-1}(\mathbf{x} - \boldsymbol{\mu}_i)\right\}$$

is the multivariate normal density; μ_i and Σ_i are mean vector and covariance matrix of i^{th} population.

In next sub-sections, a brief introduction about Bi-Normal ROC and MROC models are given.

Bi-Normal ROC model²

Let S denote the test scores in Normal/Healthy (H) and Abnormal/Diseased (D) populations, respectively,

i.e.,
$$S_H \sim N(\mu_H, \sigma_H^2)$$
, $S_D \sim N(\mu_D, \sigma_D^2)$

where μ_H , μ_D and σ_H^2 , σ_D^2 are the means and variances of H and D populations, respectively. It is assumed that the mean of population D is greater than the mean of population H (*i.e.*, $\mu_D > \mu_H$), but no constraints are placed on the standard deviations.

The false positive rate (FPR) and true positive rate (TPR) are defined as²⁰

$$FPR = P(S > c|H) = \Phi\left(\frac{\mu_H - c}{\sigma_H}\right);$$
$$TPR = P(S > c|D) = \Phi\left(\frac{\mu_D - c}{\sigma_D}\right)$$

Using FPR, we can obtain the expression of 'c' as $c=\mu_H-\sigma_H \Phi^{-1}$ [FPR], where Φ^{-1} is the inverse cumulative distribution function of Normal distribution. Then on substituting the 'c' in TPR, the expression for the Bi-normal ROC model will be

$$ROC = \Phi \left[\frac{\mu_D - \mu_H}{\sigma_D} + \frac{\sigma_H}{\sigma_D} \Phi^{-1} \left[FPR \right] \right]$$

Area under the Curve (AUC) is an important summary measure of an ROC curve and plays a prominent role in assessing the performance of a test, which takes value in the interval [0, 1]. A perfect diagnostic test is one with an area equal to 1 and a test with an area less than or equal to 0.5 is said to be inaccurate. The AUC of an ROC curve can be interpreted as the average sensitivity for all possible values of specificity and vice versa. AUC can be obtained by integrating the ROC expression over the range [0, 1]

The AUC expression for the Bi-Normal ROC is

$$AUC = P(S_D > S_H) = \Phi\left(\frac{\mu_D - \mu_H}{\sqrt{\sigma_D^2 + \sigma_H^2}}\right)$$

Multivariate ROC (MROC) model¹⁵

Let H and D be two populations assumed to follow multivariate normal distribution with mean vectors $\boldsymbol{\mu}_{H}$, $\boldsymbol{\mu}_{D}$ and covariance matrices $\boldsymbol{\Sigma}_{H}$, $\boldsymbol{\Sigma}_{D}$ respectively. The intrinsic measures are follow as:

$$FPR = 1 - \Phi\left(\frac{c - b'\boldsymbol{\mu}_H}{\sqrt{b'\boldsymbol{\Sigma}_H b}}\right); TPR = \Phi\left(\frac{b'\boldsymbol{\mu}_D - c}{\sqrt{b'\boldsymbol{\Sigma}_D b}}\right)$$

where the vector 'b' is obtained using minimax procedure²¹ as

$$b = [t \boldsymbol{\Sigma}_D + (1-t)\boldsymbol{\Sigma}_H]^{-1} (\boldsymbol{\mu}_D - \boldsymbol{\mu}_H) \quad (1)$$

here 't' is constant determined by the trial and error method in the interval (0,1) and the expression of cut-off point 'c' at each 't' can be obtain using minimax procedure by equating TPR and FPR (for detailed note on minimax approach in MROC, refer¹⁵). Upon simplifying the FPR expression and substituting the quantity c in TPR we obtain the following MROC expression.

$$MROC = \Phi\left[\frac{b'(\boldsymbol{\mu}_D - \boldsymbol{\mu}_H) - \sqrt{b'\boldsymbol{\Sigma}_H b} \, \Phi^{-1}(1 - FPR)}{\sqrt{b'\boldsymbol{\Sigma}_D b}}\right]$$

The AUC expression for the MROC is

$$AUC = P(S_D > S_H) = \Phi\left[\frac{b'(\boldsymbol{\mu}_D - \boldsymbol{\mu}_H)}{\sqrt{b'(\boldsymbol{\Sigma}_D + \boldsymbol{\Sigma}_H)b}}\right]$$

The linear combination for the MROC model is given by

$$U = b_1 X_1 + b_2 X_2 + \dots + b_k X_k \Longrightarrow U = b'X$$

where $b' = [b_p, b_2, ..., b_k]$, X is a $(k \times n)$ data matrix, k is the number of variables and here, b is the vector of coefficients obtained through equation.¹ Here, U contains the scores derived from each profile of the individuals. If $U \le c$, then the individual will be classified into H populations, otherwise into D population.

Initial works on extending two-class problem to multi-class was by.²²⁻²⁴ Further, a straight forward approach in generalizing the AUC as Volume Under Surface (VUS) was proposed by²⁵ and later some works on estimations of VUS was by.^{26, 27, 28} Even though, the articles referred above deal with both binary and multiclass problems, but their complexity will be in visualizing the ROC curve with more than 3 classes and also the computation of AUC will be bit complicated.

In this work, we have made an attempt to propose the multi-class ROC models using the flexible approach of finite mixture models. The advantages that we have noticed in working with mixture ROC models are (i) even if there are p-components in the data, still the overall ROC curve can be presented in a 2-dimensional space (ii) the ROC mixtures will provide an easy understanding to a practitioner

about classification/allocation of subjects (iii) the interpretation and presentation of summary measures will be with more ease. Hence, we have chosen EM algorithm and FMM as the mathematical treatment to determine the number of classes with in the data and also to model the mixture ROC form. For illustration purpose, the proposed methodology is given for three-class (k=3) population (figure (2)).

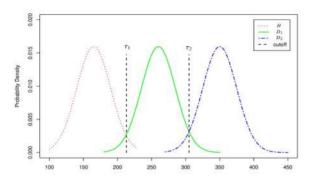


Figure 2. Hypothetically overlapping density curves and the cutoff values

The plot depict the hypothetically overlying densities of three populations and two possible cutoff points.

The subsequent sections will be on the proposed methodology, discussions with necessary illustrations using real and simulated data sets.

Proposed Methodology Mixture of Univariate Receiver Operating Characteristic (mUROC) curve

Let S denote the test scores of the identified classes as H, D_1 and D_2 populations

i.e.,
$$S_H \sim N(\mu_H, \sigma_H^2); S_{D_1} \sim N(\mu_{D_1}, \sigma_{D_1}^2)$$
 and

 $S_{D_2} \sim N(\mu_{D_2}, \sigma_{D_2}^2).$

Intrinsic measures of mUROC curve The expression for the FPR in the mixture form is defined as $FPR = \lambda_1 FPR_1 + \lambda_2 FPR_2$ (2) here, λ_1 and λ_2 are the mixing proportions and will be obtained using EM algorithm. By definition, FPR_1 and FPR_2 are derived from H and D_1 populations, and will take the following forms

$$FPR_1 = \Phi\left(\frac{\mu_H - c_1}{\sigma_H}\right); FPR_2 = \Phi\left(\frac{\mu_{D_1} - c_2}{\sigma_{D_1}}\right)(3)$$

using equation (3), c_1 and c_2 can be written as

$$c_1 = \mu_H - \sigma_H \Phi^{-1} (FPR_1); c_2 = \mu_{D1} - \sigma_{D1} \Phi^{-1} (FPR_2)$$
(4)

where Φ^{-1} is the inverse cumulative distribution function of Normal distribution. By definition, the TPR₁ and TPR₂ are obtained from D₁ and D₂ populations, and are as follows

$$TPR_{1} = \Phi\left(\frac{\mu_{D_{1}} - c_{1}}{\sigma_{D_{1}}}\right); TPR_{2} = \Phi\left(\frac{\mu_{D_{2}} - c_{2}}{\sigma_{D_{2}}}\right)$$
(5)

The expression for the TPR in the mixture form is defined as

$$TPR = \lambda_1 TPR_1 + \lambda_2 TPR_2 \tag{6}$$

by substituting equation (4) in equation (5) and thereafter equation (5) in equation (6), we get the expression for mixture ROC model

$$mUROC = \lambda_1 \left[\Phi(A_1 + B_1 \Phi^{-1} (FPR_1)) + \lambda_2 \right]$$
$$\left[\Phi(A_2 + B_2 \Phi^{-1} (FPR_2)) \right]$$
(7)

which is called a 2-component mixture of Univariate Normal Receiver Operating Characteristic (mUROC) curve, where

$$A_1 = \frac{\mu_{D_1} - \mu_H}{\sigma_{D_1}}, A_2 = \frac{\mu_{D_2} - \mu_{D_1}}{\sigma_{D_2}}, B_1 = \frac{\sigma_H}{\sigma_{D_1}}$$

 $B_2 = \frac{\sigma_{D_1}}{\sigma_{D_2}}$. The AUC expression for mUROC curve will be

$$mUAUC = \lambda_1 \Phi\left(\frac{A_1}{\sqrt{1+B_1^2}}\right) + \lambda_2 \Phi\left(\frac{A_2}{\sqrt{1+B_2^2}}\right)$$
(8)

In general, the p-component mixture ROC expression of equation (7) and AUC of equation (8) can be expressed as

$$mUROC_{p} = \sum_{i=1}^{p-1} \lambda_{i} \left[\Phi(A_{i} + B_{i} \Phi^{-1}(FPR_{i})) \right]$$
$$mUAUC_{p} = \sum_{i=1}^{p-1} \lambda_{i} \Phi\left(\frac{A_{i}}{\sqrt{1 + B_{i}^{2}}}\right)$$

where

$$\sum_{i=1}^{p-1} \lambda_i = 1; A_i = \frac{\mu_{i+1} - \mu_i}{\sigma_{i+1}}; B_i = \frac{\sigma_i}{\sigma_{i+1}}; i = 1, 2, \dots, (p-1)$$

If there are 'p' populations, there will be (p-1) linear combinations. Hence, the summation part λ_i 's and ROC expression will range from 1 to (p-1).

Mixture of Multivariate Receiver Operating Characteristic (mMROC) Curve

Let $S_{H} \sim N_{p} (\boldsymbol{\mu}_{H}, \boldsymbol{\Sigma}_{H})$; $S_{D_{1}} \sim N_{p} (\boldsymbol{\mu}_{D_{1}}, \boldsymbol{\Sigma}_{D_{1}})$ and $S_{D_{2}} \sim Np (\boldsymbol{\mu}_{D_{2}}, \boldsymbol{\Sigma}_{D_{2}})$, where H, D_{1}

and D_2 are class labels pertaining to healthy/ normal (H) and two abnormal populations, namely D_1 and D_2 .

Intrinsic measures of mMROC curve

The expression for FPR in the mixture form is defined as

$$FPR = \lambda_1 \ FPR_1 + \lambda_2 \ FPR_2$$
(9)

here

$$FPR_{1} = 1 - \Phi\left(\frac{c_{1} - b_{1}'\boldsymbol{\mu}_{H}}{\sqrt{b_{1}'\boldsymbol{\Sigma}_{H} b_{1}}}\right); FPR_{2} = 1 - \Phi\left(\frac{c_{2} - b_{2}'\boldsymbol{\mu}_{D_{1}}}{\sqrt{b_{2}'\boldsymbol{\Sigma}_{D_{1}} b_{2}}}\right)$$

where b_1 , and $b_2 (\neq 0)$ are the vector of coefficients of k variables obtained from a minimax procedure and are given as

$$\begin{array}{l} b_1 = [t\Sigma_{D1} + (1-t) \quad \Sigma_H]^{-1} \quad (\mu_{D1} - \mu_H); b_2 = [t\Sigma_{D_2} + (1-t) \\ \Sigma_{D_1}]^{-1} \quad (\mu_{D_2} - \mu_{D_1}) \end{array}$$

where 't' is a constant lies between 0 and 1 and determined by trial and error method with an increment of 0.1. Here, b_1 corresponds to vector of coefficients of H and D_1 populations and b_2 corresponds to vector of coefficients of D_1 and D_2 populations. From equation (10), c_1 and c_2 can be written as

$$c_{1} = b_{1}^{\prime} \boldsymbol{\mu}_{H} + \sqrt{b_{1}^{\prime} \boldsymbol{\Sigma}_{H} b_{1}} \Phi^{-1} (1 - FPR_{1});$$

$$c_{2} = b_{2}^{\prime} \boldsymbol{\mu}_{D_{1}} + \sqrt{b_{2}^{\prime} \boldsymbol{\Sigma}_{D_{1}} b_{2}} \Phi^{-1} (1 - FPR_{2})$$

where Φ^{-1} is the inverse cumulative distribution of multivariate normal. The expression for TPR in the mixture form is defined as

$$TPR = \lambda_1 TPR_1 + \lambda_2 TPR_2$$
(11)

Where;

$$TPR_{1} = \Phi\left(\frac{b_{1}'\boldsymbol{\mu}_{D_{1}} - c_{1}}{\sqrt{b_{1}'\boldsymbol{\Sigma}_{D_{1}}b_{1}}}\right); TPR_{2} = \Phi\left(\frac{b_{2}'\boldsymbol{\mu}_{D_{2}} - c_{2}}{\sqrt{b_{2}'\boldsymbol{\Sigma}_{D_{2}}b_{2}}}\right)$$

Substituting the expressions of c_1 and c_2 in equation (11), we obtain the ROC expression

in the following mixture form as

mMROC =
$$\lambda_1 \Phi[\alpha_1 + \beta_1 \Phi^{-1} (1-FPR_1)]$$

+ $\lambda_2 \Phi[\alpha_2 + \beta_2 \Phi^{-1} (1-FPR_2)]$ (12)

which is called a 2- component mixture of Multivariate Receiver Operating Characteristic (mMROC) curve where

$$\alpha_{1} = \frac{b_{1}'(\mu_{D_{1}} - \mu_{H})}{\sqrt{b_{1}'\Sigma_{D_{1}}b_{1}}}, \ \alpha_{2} = \frac{b_{2}'(\mu_{D_{2}} - \mu_{D_{1}})}{\sqrt{b_{2}'\Sigma_{D_{2}}b_{2}}},$$
$$\beta_{1} = -\frac{\sqrt{b_{1}'\Sigma_{H}b_{1}}}{\sqrt{b_{1}'\Sigma_{D_{1}}b_{1}}} \beta_{2} = -\frac{\sqrt{b_{2}'\Sigma_{D_{1}}b_{2}}}{\sqrt{b_{2}'\Sigma_{D_{2}}b_{2}}}.$$

The AUC expression for mMROC curve will be

$$mMAUC = \lambda_1 \Phi\left(\frac{\alpha_1}{\sqrt{1+\beta_1^2}}\right) + \lambda_2 \Phi\left(\frac{\alpha_2}{\sqrt{1+\beta_2^2}}\right)$$
(13)

In general, a p-component mixture ROC expression of equation (12) and AUC of equation (13) can be expressed as

$$mMROC_{p} = \sum_{i=1}^{p-1} \lambda_{i} \Phi[\alpha_{i} + \beta_{i} \Phi^{-1}(1 - FPR_{i})]$$
$$mMAUC_{p} = \sum_{i=1}^{p-1} \lambda_{i} \Phi\left(\frac{\alpha_{i}}{\sqrt{1 + \beta_{i}^{2}}}\right)$$

Here

$$\sum_{i=1}^{p-1} \lambda_i = 1, \, \alpha_i = \frac{b'_i (\mu_{i+1} - \mu_i)}{\sqrt{b'_i \Sigma_{i+1} b_i}},$$
$$\beta_i = -\frac{\sqrt{b'_i \Sigma_i b_i}}{\sqrt{b'_i \Sigma_{i+1} b_i}}, i = 1, 2, \dots, (p-1)$$

Youden's Index

Usually, the test's capability or the performance of a biomarker depends on the optimal threshold, which provides maximum degree of correct classification. In order to determine the optimal threshold Youden's Index (J) is most preferably used measure and used by many researchers in the context to determine the optimal threshold,²⁹⁻³² it is given as

J = Max (TPR + TNR - 1)

Here, TNR is the true negative rate (TNR = 1-FPR). Using the expressions of mUROC curve and mMROC curve, equations (2, 6, 9 & 11), Youden's index takes the following forms

$$J_{U} = Max \left\{ \lambda_{1} \left[\Phi\left(\frac{\mu_{D_{1}} - c_{1}}{\sigma_{D_{1}}}\right) + \Phi\left(\frac{c_{1} - \mu_{H}}{\sigma_{H}}\right) \right] + \lambda_{2} \left[\Phi\left(\frac{\mu_{D_{2}} - c_{2}}{\sigma_{D_{2}}}\right) + \Phi\left(\frac{c_{2} - \mu_{D_{1}}}{\sigma_{D_{1}}}\right) \right] \right\} (14)$$

$$J_{M} = Max \left\{ \lambda_{1} \left[\Phi\left(\frac{b_{1}' \mu_{D_{1}} - c_{1}}{\sqrt{b_{1}' \Sigma_{D_{1}} b_{1}}} \right) + \Phi\left(\frac{c_{1} - b_{1}' \mu_{H}}{\sqrt{b_{1}' \Sigma_{H} b_{1}}} \right) \right] + \lambda_{2} \left[\Phi\left(\frac{b_{2}' \mu_{D_{2}} - c_{2}}{\sqrt{b_{2}' \Sigma_{D_{2}} b_{2}}} \right) + \Phi\left(\frac{c_{2} - b_{2}' \mu_{D_{1}}}{\sqrt{b_{2}' \Sigma_{D_{1}} b_{2}}} \right) \right] \right\} (15)$$

where $J_U \& J_M$ are the Youden's indices of univariate and multivariate mixture ROC curves. Using the equations (14 & 15), the corresponding score related to the max J_U and J_M will be taken as optimal thresholds, say τ_1 and τ_2 , for classifying into several populations.

Comparing two AUCs

In order to compare the AUC's obtained from the different models under univariate and multivariate setup, the null and alternative hypotheses along with the test statistic is given for both univariate and multivariate cases respectively.

For univariate case:

 $H_0: AUC = mUAUC \ Vs \ H_1: AUC \neq mUAUC$

$$Z_{AUC} = \frac{\widehat{AUC} - \widehat{mUAUC}}{\sqrt{var(\widehat{AUC}) + var(\widehat{mUAUC})}}$$

For multivariate case:

$$H_0: AUC = mMAUC \ Vs \ H_1: AUC \neq mMAUC$$

$$Z_{AUC} = \frac{\widehat{AUC} - m\widehat{MAUC}}{\sqrt{var(\widehat{AUC}) + var(m\widehat{MAUC})}}$$

The test statistic value follows normal distribution with significance level α asymptotically.

Table 1. ROC Curve parameters and AUC measures
--

Results and Discussions

The proposed methodology is supported with real data sets, namely, OGTT¹⁶ for univariate and Disk Hernia³³ for multivariate ROC models, respectively. Simulations are performed for mUROC, mMROC, Bi-Normal ROC and MROC curves, respectively with different parameter combinations at varying sample sizes. Detailed discussion along with necessary results are presented in different sub sections.

Results of mUROC model Real data set

The OGTT data (n=21) is visualized using histogram and an overlying density curve in figure (1). Figure 1(a) depicts the tri-model pattern of OGTT data and on applying EM-algorithm, 3-components are witnessed (see figure 1(b)). The proposed mUROC model is then applied to this 3-component data. The estimated summary and intrinsic measures cut-off are reported in table (1). The estimated mixing proportions are $\hat{\lambda}_1 = 0.4992$ and $\hat{\lambda}_2 = 0.5008$.

From the results of mUROC model, it is seen that identification of a hidden population helped out in exhibiting the true accuracy and

D: Normal	$\mu_{\rm H}$	h	D	$\sigma_{_{\rm H}}$	σ	D		c	FPR	TPR	AUC	J	Z _{AUC}	Sig
Bi-Normal	6.7118	15.5	5550	1.9840	6.4	567	9.	79	0.0603	0.8140	0.9048	0.7537		
mUROC	$\mu_{\rm H}$	μ_{D_1}	$\boldsymbol{\mu}_{D_2}$	σ _н	σ _{D1}	σ _{D2}	τ ₁	τ ₂	FPR	TPR	mUAUC	$J_{_{\rm U}}$	1.99932	0.02279*
	6.7118	9.8793	21.2127	1.9840	1.6068	2.8680	9.22	12.06	0.1013	0.8273	0.9463	0.7260		

*Significant

The table consists of parameters of conventional Bi-Normal ROC and the proposed mixture univariate ROC model and its AUC measures, respectively.

reliable information in the data and which was at higher value than the Bi-Normal ROC model. This result and claim is based on using the expressions given in the above section of comparing two AUC's. From the results of Z statistic, it is clearly evident that the proposed mixture univariate ROC model provides better accuracy by accounting the hidden components in the model than the Bi-Normal ROC which has only two components. So, even if you have known class labels, always it is better to investigate further for observing hidden sub components.

The optimal cut points τ_1 (9.22) and τ_2 (12.06) are obtained at max $J_U = 0.7260$. In similar manner, the optimal cut point for the Bi-Normal ROC model is observed at max J =0.7537 i.e., c = 9.79. If the test score is greater than the optimal cut point $\tau_2 = 12.06$, then the individual is classified to diseased component 2 (D₂) else if the score lies between $\tau_1 = 9.22$ and $\tau_2 = 12.06$, then the individual is classified to diseased component 1 (D₁) else individual will be assigned in healthy population (H). The above explanation can be presented in the following way

The individual is classified as = $\begin{cases}
Normal & if \quad S \leq \tau_1 \\
Diseased \ Component \ 1 & if \quad \tau_1 \leq S \leq \tau_2 \\
Diseased \ Component \ 2 & if \quad S > \tau_2
\end{cases}$

The ROC curves for Bi-Normal and mUROC are shown in figure (3). From the graph it can

be clearly understood that mUROC curve is slightly superior with FPR (0.1013) and higher TPR (0.8273) values than that of Bi-Normal ROC curve (0.0603, 0.8140).

Simulation studies

In table (2), four sets of means and variances are considered along with the initial values for mixing proportions. Of which, the first two sets (A & B) are with unequal variances and the last two sets (C & D) are with equal variances. In each population, random samples of size n = $\{25,50,100,250,500\}$ were generated using the parameter values defined in four sets and for each set, 1000 iterations were performed.

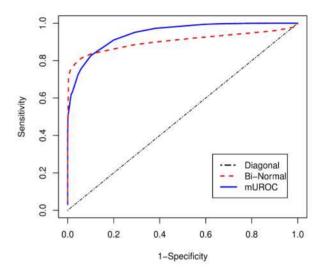


Figure 3. ROC Curves for the OGTT data estimated from Bi-Normal and mUROC

Sets	$\frac{\delta_{1}}{\lambda_{1}}$	λ_2	$\mu_{\rm H}$	μ_{D_1}	μ_{D_2}	$\sigma_{\rm H}$	σ_{D_1}	σ_{D_2}
А	0.5	0.5	30.3	33.5	42.2	1.0	1.5	2.0
В	0.5	0.5	30.3	30.3	30.3	1.0	1.5	2.0
С	0.5	0.5	30.3	33.5	42.2	1.5	1.5	1.5
D	0.5	0.5	30.3	30.3	30.3	1.5	1.5	1.5

 Table 2. Mixing Proportions, Means and Variances are considered for simulation studies

The table depicts the four sets of means and variances where the first two sets have unequal variances while the other have equal variances.

Table 3. Bi-Normal ROC and mUROC parameters at various sample sizes

	Sample	Bi-Normal ROC parameters					mUROC parameters							
Set	size	$\hat{\mu}_{H}$	$\hat{\mu}_{D}$	$\hat{\sigma}_{H}$	$\hat{\sigma}_D$	$\hat{\lambda}_1$	$\hat{\lambda}_2$	$\hat{\mu}_{_{\rm H}}$	$\hat{\mu}_{D_1}$	$\hat{\mu}_{D_2}$	$\hat{\sigma}_{H}$	$\hat{\sigma}_{D_1}$	$\hat{\sigma}_{D_2}$	
	25	30.28999	37.83501	0.99371	4.80298	0.49897	0.50104	30.28999	33.53889	42.24360	0.99371	1.42354	2.28759	
	50	30.29443	37.80698	0.99815	4.80505	0.49487	0.50513	30.29443	33.52135	42.19013	0.99815	1.44216	2.43348	
А	100	30.29929	37.84866	0.99618	4.80581	0.49794	0.50207	30.29929	33.50459	42.19767	0.99618	1.47008	2.47394	
	250	30.30241	37.85816	0.99984	4.81676	0.49865	0.50135	30.30241	33.50111	42.19369	0.99984	1.48849	2.47896	
	500	30.30145	37.83916	1.00027	4.80862	0.50008	0.49992	30.30145	33.50060	42.20033	1.00027	1.49460	2.48484	
	25	30.30087	30.27936	0.98916	2.02794	0.50893	0.49107	30.30087	30.30915	30.32666	0.98916	1.46140	2.34095	
	50	30.29896	30.29042	0.99306	2.04269	0.50735	0.49265	30.29896	30.29435	30.30553	0.99306	1.47204	2.41244	
В	100	30.30428	30.30635	0.99711	2.05678	0.50247	0.49753	30.30428	30.29453	30.28683	0.99711	1.49322	2.45562	
	250	30.30187	30.29860	0.99877	2.05799	0.50106	0.49894	30.30187	30.29931	30.29915	0.99877	1.49452	2.48817	
	500	30.30071	30.30195	0.99981	2.06128	0.50039	0.49961	30.30071	30.29562	30.29857	0.99981	1.49729	2.49579	
	25	30.29027	37.86248	1.47544	4.58276	0.49819	0.50181	30.29027	33.52860	42.19653	1.47544	1.38771	1.40285	
	50	30.30231	37.87722	1.48463	4.59170	0.50002	0.49998	30.30231	33.51888	42.20491	1.48463	1.47319	1.44567	
С	100	30.29901	37.86276	1.49659	4.59827	0.50170	0.49830	30.29901	33.50236	42.19209	1.49659	1.47012	1.47389	
	250	30.30195	37.85228	1.49666	4.59508	0.50047	0.49953	30.30195	33.50322	42.19834	1.49666	1.49416	1.49620	
	500	30.29718	37.84655	1.49739	4.60281	0.50041	0.49959	30.29718	33.50430	42.19980	1.49739	1.49373	1.49321	
	25	30.29557	30.28307	1.48423	1.49149	0.50246	0.49754	30.29557	30.30278	30.30278	1.48423	1.45850	1.45850	
	50	30.30643	30.30014	1.49269	1.48801	0.50013	0.49987	30.30643	30.31424	30.31424	1.49269	1.47716	1.47716	
D	100	30.30320	30.30215	1.49595	1.49531	0.50345	0.49655	30.30320	30.30023	30.30023	1.49595	1.49175	1.49175	
	250	30.30292	30.30036	1.49844	1.49698	0.50925	0.49075	30.30292	30.30347	30.30347	1.49844	1.49467	1.49467	
	500	30.29990	30.29905	1.50123	1.49692	0.50345	0.49655	30.29990	30.30083	30.30083	1.50123	1.49714	1.49714	

The table depicts the estimates of mixing proportions, means and variances of the four sets considered for simulation studies at various sample sizes.

Using EM algorithm, mixing proportions, means and variances were estimated (table, 3) and table (4) has the estimates of the intrinsic and AUC measures of Bi-Normal ROC and mUROC curves at different sample sizes.

In continuation to the results obtained through

simulation studies presented in tables (3 & 4), table (5) is about the AUC comparisons between Bi-Normal and mUROC models respectively. The p-value at each sample size for respective sets A and C indicate that there is evidence that the AUC obtained through

	Sample	Е	i-Normal R	OC measure	s		mUI	ROC measur	res	
Set	size	с	FPR	TPR	J	τ_1	τ_2	FPR	TPR	JU
	25	37.66963	0.02022	0.85858	0.83836	30.33720	37.98989	0.00808	0.93907	0.93099
	50	37.35715	0.00684	0.86308	0.85623	30.41233	37.71593	0.00406	0.93922	0.93516
А	100	37.53743	0.01767	0.86980	0.85213	30.37432	37.85399	0.00666	0.94144	0.93478
	250	37.33082	0.01682	0.87378	0.85696	30.37836	38.00560	0.00778	0.94628	0.93849
	500	37.33457	0.02269	0.88892	0.86623	30.39347	37.55582	0.00892	0.95067	0.94175
	25	30.27284	0.10281	0.21574	0.11293	30.34705	30.43388	0.08364	0.24271	0.15908
	50	30.37245	0.08438	0.22962	0.14523	30.35976	30.38076	0.08964	0.24785	0.15821
В	100	30.36387	0.09169	0.23242	0.14073	30.39627	30.47292	0.10124	0.25127	0.15002
	250	30.42390	0.09125	0.23347	0.14221	30.37676	30.44896	0.07135	0.25767	0.18632
	500	30.48969	0.08408	0.25025	0.16617	30.39324	30.54108	0.08038	0.26974	0.18936
	25	37.48985	0.04690	0.82295	0.77605	30.38452	37.84553	0.02949	0.88754	0.85804
	50	37.64905	0.03460	0.82665	0.79206	30.40888	37.85051	0.02231	0.89500	0.87269
С	100	37.43438	0.03144	0.86252	0.83108	30.39531	38.02926	0.02631	0.89541	0.86909
	250	37.40727	0.02916	0.86581	0.83665	30.36950	37.65668	0.02459	0.90736	0.88276
	500	37.37727	0.03213	0.86587	0.83374	30.45430	38.01432	0.03082	0.90992	0.87911
	25	30.25927	0.10035	0.11446	0.01411	30.32118	30.23058	0.17075	0.20607	0.03531
	50	30.39390	0.16276	0.22156	0.05880	30.33703	30.34987	0.18308	0.20898	0.02590
D	100	30.27969	0.77375	0.79293	0.01918	30.40987	30.32450	0.23867	0.24340	0.00474
	250	30.32154	0.78966	0.80224	0.01258	30.29218	30.31667	0.82486	0.83294	0.00808
	500	30.30986	0.90768	0.91334	0.00566	30.24286	30.32719	0.88734	0.89423	0.00690

T 11 4 D'NT	1 DOO	1 1000			
Table 4. Bi-Norma	l ROC and	1 mL/R()()	measures af	Various s	ample sizes
		amonoo	measures at	various s	

The table depicts the measures of Bi-Normal and mUROC model of the four sets considered for simulation studies at various sample sizes.

		e		•	
Set	Sample size	ÂUC	mÛÂÛC	Z _{AUC}	Sig.
	25	0.93479	0.97863	1.65738	0.04872*
	50	0.93544	0.98012	2.07329	0.01907*
А	100	0.93714	0.98013	2.88499	0.00196*
	250	0.93736	0.98035	4.47766	0.00000*
	500	0.93741	0.98035	6.43528	0.00000*
	25	0.49611	0.50264	0.06851	0.47269 ^{NS}
	50	0.49869	0.49942	0.05286	0.47892^{NS}
В	100	0.49856	0.50045	0.28207	0.61105 ^{NS}
	250	0.49943	0.49995	0.03249	0.48704^{NS}
	500	0.49943	0.50021	0.03203	0.48722 ^{NS}
	25	0.93773	0.96783	1.64521	0.04996*
	50	0.93973	0.96694	1.85489	0.03181*
С	100	0.94013	0.96693	1.94829	0.02569*
	250	0.94052	0.96718	2.63752	0.00418*
	500	0.94040	0.96721	3.68464	0.00011*
	25	0.49716	0.50026	0.03908	0.48441 ^{NS}
D	50	0.49883	0.50098	0.02551	0.48982 ^{NS}
	100	0.49970	0.49996	0.01895	0.49244 ^{NS}
	250	0.49953	0.50000	0.01998	0.49203 ^{NS}
	500	0.49983	0.50023	0.01656	0.49340 ^{NS}

Table 5. The estimated AUC's of Bi-Nomral and mUROC along with values at various sample sizes

*Significant;

NS, Not significant.

mUROC model is comparatively better than the existing Bi-Normal ROC model. The sets A & C are taken in such a way that they exhibited the better classification and sets B & D relates to explain the behaviour of worst classification scenario. Since the sets B & D are taken to mimic the worst cases situation, the insignificant p-values are witnessed with nearer mean values of Bi-Normal and mUROC models.

The Bi-Normal and mUROC curves for each set at various sample sizes can be seen in figure (4) & (5). It is observed that, in each graph, the conventional Bi-Normal and proposed Mixture

ROC curves almost overlap each other and which indicates that the extent of classification is similar at varying sample sizes. Here, the primary point is on focusing the use of mUROC curve, when unseen or hidden populations are extracted. In such cases, the Bi-Normal ROC may not be used and masks the true accuracy, FPR and TPR.

Results of mMROC model Real data set

In multivariate case, to demonstrate the practical applicability of the proposed mMROC

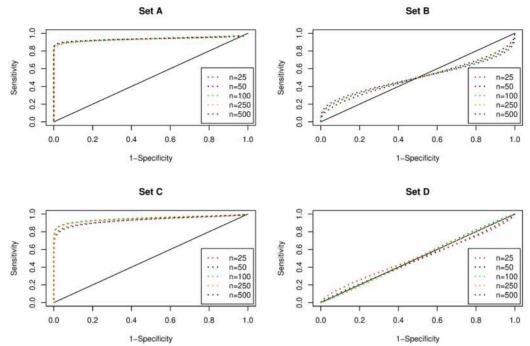


Figure 4. Bi-Normal ROC Curves for the simulated data sets at various sample sizes The Bi-Normal ROC curves for the four sets are shown here. The first and third sets are the examples of best cases while the second and fourth are the worst cases of classification.

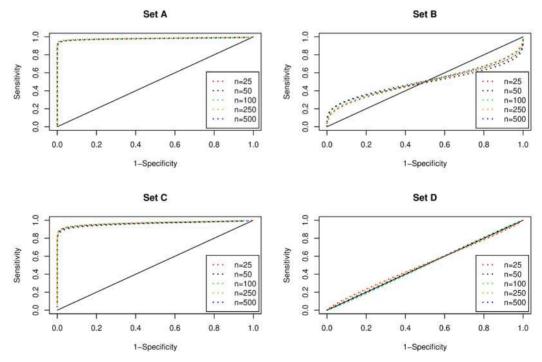


Figure 5. mUROC Curves for the simulated data sets at various sample sizes

The mUROC curves for the four sets are shown here. The first and third sets are the examples of best cases while the second and fourth are the worst cases of classification.

Multi-Class Classification using Mixtures of Univariate and ...

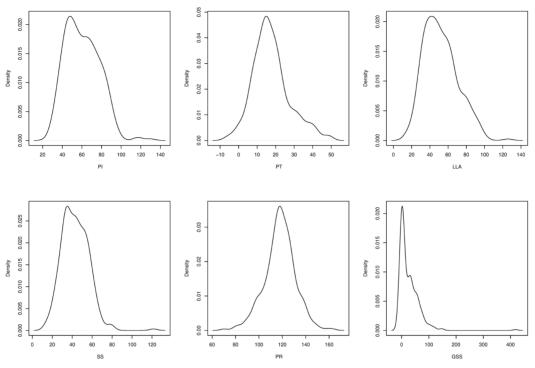


Figure 6. Density plots of each variable in Disk Hernia data set Density plots of each variable in Disk Hernia data set shown here for observing the density patterns of each variable.

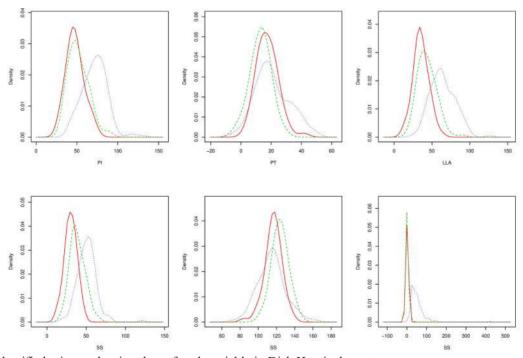


Figure 7. Identified mixture density plots of each variable in Disk Hernia data set The identified components of each variable in Disk Hernia data set shown here, results shows that there are tri-modal patters exists in the data.

model, Disk Hernia data set is considered. It consists of 310 samples and 6 variables (PI, PT, LLA, SS, PR, GS) the density plots are for each variable is presented in figure (6).

Actually, this data set has three categories: Normal (N), Spondylolisthesis (ST) and Disk Hernia (DH). But, for illustration purpose, these class labels were ignored and on the complete data, EM algorithm has been applied to expose and extract the hidden components.

MROC Model

Upon this exercise, three components were identified and depicted in figure (7).

The estimated mean vectors and covariance matrices of MROC and mMROC models are

MROC Model

(^{64.6921})

Using these mean vectors and covariance matrices, the vectors b, b_1 and b_2 values, the intrinsic measures and summary

	Ĥ	$\hat{u}_{H} = \begin{pmatrix} 43.5\\ 38.8\\ 123.8\\ 2.18 \end{pmatrix}$	638 3912	$u_D = \begin{cases} 44.9 \\ 115.0 \end{cases}$	252 015 0774 776	
	/152.9650	53.1515	107.1944	99.8139	-54.6642	15.6189
	53.1515	45.9502	25.3629	7.1993	-21.1770	7.7775
γ-	107.1944	25.3629	152.8087	81.8374	-27.5706	20.3356
$\Delta_H =$	99.8139	7.1993	81.8374	92.6171	-33.4884	7.8483
	-54.6642	-21.1770	-27.5706	-33.4884	81.2478	-3.3084
	\ 15.6189	7.7775	20.3356	7.8483	-3.3084	39.7785 [/]
	/311.9394	105.9087	236.1581	206.0250	-20.9393	453.9016 \
	105.9087	110.5790	78.8221	-4.6700	36.3767	137.0192
γ	236.1581	78.8221	386.8685	157.3297	19.0751	397.3477
$z_D =$	206.0250	-4.6700	157.3297	210.6891	-57.3146	316.8820
	-20.9393	36.3767	19.0751	-57.3146	198.5553	83.9594
	\453.9016	137.0192	397.3477	316.8820	83.9594	1656.2245 [/]

 $(\begin{array}{c} 51.6856 \\ 12.8218 \end{array})$

mMROC model

$\hat{\lambda}_1 = 0.7171$, $\hat{\lambda}_2 = 0.2829$

β,	$_{H} = \begin{pmatrix} 51.68\\12.82\\43.54\\38.86\\123.89\\2.182 \end{pmatrix}$	$ \begin{array}{c} 18\\23\\38\\912 \end{array}, \hat{\mu}_{l} $	$p_1 = \begin{pmatrix} 58.12\\15.32\\49.00\\42.79\\112.3\\22.48 \end{pmatrix}$	$\begin{pmatrix} 218 \\ 018 \\ 924 \\ 824 \end{pmatrix}, \hat{\mu}$	$\hat{a}_{D_2} = \begin{pmatrix} 31.1\\73.4\\50.2\\121.2 \end{pmatrix}$	8650 183 723 2471 9078
â	$\begin{pmatrix} 152.9650\\53.1515\\107.1944 \end{pmatrix}$	53.1515 45.9502 25.3629	107.1944 25.3629 152.8087	99.8139 7.1993 81.8374	-54.6642 -21.1770 -27.5706	15.6189 7.7775 20.3356
$\Sigma_H =$	99.8139	7.1993	81.8374	92.6171	-33.4884	7.8483
	-54.6642	-21.1770	-27.5706	-33.4884	81.2478	-3.3084
	15.6189	7.7775	20.3356	7.8483	-3.3084	39.7785
$\hat{\Sigma}_{D_1} =$	214.9352	42.7037	186.1923	172.2225	-56.5213	178.1567
	42.7037	48.6219	9.7010	-5.9223	8.5664	-19.1297
	186.1923	9.7010	264.8324	176.4905	-63.2351	218.4417
	172.2225	-5.9223	176.4905	178.1399	-65.0804	197.2867
	-56.5213	8.5664	-63.2351	-65.0804	171.3358	-69.6850
	178.1567	-19.1297	218.4417	197.2867	-69.6850	371.3049

$\hat{\Sigma}_{D_2} =$	$\begin{pmatrix} 164.8756\\ 0.9426\\ 49.1751\\ 163.9377\\ -89.2222 \end{pmatrix}$	0.9426 86.8138 -24.5043 -85.8586 -1.6486	49.1751 -24.5043 260.2649 -24.6890 60.2220	163.9377 -85.8586 -24.6890 249.7885 -87.5861	-89.2222 -1.6486 60.2220 -87.5861 199.1375	244.2145 -81.5979 -104.0618 325.8146 102.8736
	-89.2222	-1.6486	60.2220	-87.5861	199.1375	102.8736 2791.2693
	244.2145	-81.5979	-104.0618	325.8146	102.8736	2791.2693 [/]

Table 6. Cutoffs and measures of MROC curve for varying t values

t	с	FPR	TPR	AUC	J
0.1	-15.7252	0.1920	0.8080	0.8521	0.6159
0.2	-14.8812	0.1847	0.8153	0.8708	0.6306
0.3	-13.8502	0.1837	0.8163	0.8790	0.6325
0.4	-12.0972	0.1875	0.8125	0.8822	0.6250
0.5	-12.9164	0.1852	0.8148	0.8830	0.6297
0.6	-11.3782	0.1902	0.8098	0.8825	0.6196
0.7	-10.7431	0.1929	0.8071	0.8813	0.6141
0.8	-10.1777	0.1956	0.8044	0.8798	0.6088
0.9	-9.6706	0.1982	0.8018	0.8782	0.6036

The table consists of MROC measures at various values of 't', highlighting the optimal cut point identified using Youden's Index (J)

measures of MROC and mMROC curve were computed and are presented in tables (5) and (6), respectively. The optimal cut points of mMROC are $\tau_1 = -13.0535$ and $\tau_2 =$ 24.0451 are observed at max $J_M = 0.7294$ with mMAUC = 0.9324 at t = 0.5. At t = 0.5, the linear combinations for MROC model and mMROC models are

MROC model

mMROC model

U₁=-13.0695*PI+13.1404*PT-0.0291*LLA+12.9842*SS-0.0900*PR+0.1440*GS

 U_{II} =5.0773*PI-4.7765*PT+0.0784*LLA-5.0192*SS+0.0676*PR+0.0306*GS Here, U denote the scores derived from the linear combinations of MROC model. Similarly, U_I and U_{II} are the scores obtained from the linear combinations of 2-component mMROC model.

If the test score is greater than $\tau_2 = 24.0451$, then individual is classified as diseased category, namely, Disk Hernia (D_2) else if the score lies between $\tau_1 = -13.0535$ and $\tau_2 = 24.0451$, the individual is classified into diseased category, Spondylolisthesis (D_1) else individual considered to be in healthy (H) population. From the results, it is observed that the AUC of proposed mMROC model (mMAUC=0.9324) is better than the AUC of MROC curve (AUC=0.8830). This shows that mMROC has the better accuracy of correct classification (Z=2.431513, p-value=0.007518).

In figure (8), the ROC curves for MROC and mMROC are depicted at t = 0.5. This reveals the fact that the performance is superior with mMROC than the MROC curve.

t	$ au_1$	τ_2	FPR	TPR	mMAUC	J _M
0.1	-17.4021	16.1824	0.1523	0.8477	0.9062	0.6954
0.2	-15.7196	18.0715	0.1414	0.8586	0.9206	0.7172
0.3	-14.5492	19.9254	0.1367	0.8633	0.9280	0.7266
0.4	-13.6936	21.8854	0.1353	0.8647	0.9314	0.7279
0.5	-13.0535	24.0466	0.1360	0.8640	0.9324	0.7294
0.6	-12.5701	26.5148	0.1381	0.8619	0.9316	0.7237
0.7	-12.2059	29.4317	0.1412	0.8588	0.9298	0.7177
0.8	-11.9355	33.0070	0.1448	0.8552	0.9271	0.7104
0.9	-11.7414	37.5768	0.1489	0.8511	0.9240	0.7022

Table 7. Cutoffs and measures of mMROC curve for varying t values

The table consists of mMROC measures at various values of 't', highlighting the optimal cut point identified using Youden's Index (J)

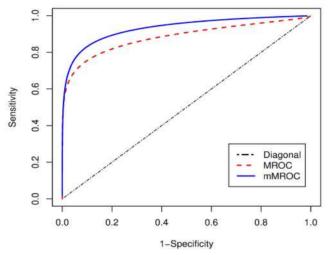


Figure 8. The MROC and mMROC curves for Disk Hernia data set

Simulation studies

Simulation studies are carried out for the proposed mMROC using bivariate normal random variables $(X_1 \text{ and } X_2)$. In table (7), four sets of mean vectors and covariance matrices were considered along with the initial mixing proportions to depict the behaviour of the proposed mMROC. Of these, the first two sets (A & B) are with unequal covariances matrices and the last two sets (C & D) are with equal covariance matrices. In each population, random samples of size n =

{25,50,100,250,500} were generated.

Table (8) has the estimated mean vectors and covariance matrices of MROC model at different sample sizes. The estimated mixing proportions, mean vectors and covariance matrices of the proposed mMROC model are reported in table (9).

Table (10) and (11) depicts the linear combinations, cut-off points and measures of MROC and mMROC curves for various sample sizes at t = 0.5.

Furthermore, the AUC comparisons for simulated studies at different sample sizes

Sets	λ_1	λ ₂	μ_{H}	_	μ_{D_1}		μ_{D_2}	Σ _H	Σ _{D1}	Σ_{D_2}
1	0.5	0.5	(10.1 6.7) (13.2 9.8) (21.8 12.1	$\left(\begin{array}{c} \Sigma_{\rm H} \\ \hline 2 & 1 \\ 1 & 3 \end{array}\right)$	$\left(\begin{array}{cc}2&1\\1&4\end{array}\right)$	$ \left(\begin{array}{cc} 2 & 1 \\ 1 & 6 \end{array}\right) $
2	0.5	0.5	$\begin{pmatrix} 10.1 \\ 6.7 \end{pmatrix}$) (10.1 6.7) (10.1) 6.7	$\left(\begin{array}{cc} 2 & 1 \\ 1 & 3 \end{array}\right)$	$\left(\begin{array}{cc} 2 & 1 \\ 1 & 4 \end{array}\right)$	$\left(\begin{array}{cc} 2 & 1 \\ 1 & 6 \end{array}\right)$
3	0.5	0.5	$\begin{pmatrix} 10.1 \\ 6.7 \end{pmatrix}$) (13.2 9.8) (21.8 12.1	$\left(\begin{array}{cc} 2 & 1 \\ 1 & 4 \end{array}\right)$	$\left(\begin{array}{cc} 2 & 1 \\ 1 & 4 \end{array}\right)$	$\left(\begin{array}{cc} 2 & 1 \\ 1 & 4 \end{array}\right)$
4	0.5	0.5	$\begin{pmatrix} 10.1 \\ 6.7 \end{pmatrix}$) (10.1 6.7) (10.1) 6.7	$\left(\begin{array}{cc} 2 & 1 \\ 1 & 4 \end{array}\right)$	$\left(\begin{array}{cc} 2 & 1 \\ 1 & 4 \end{array}\right)$	$\left(\begin{array}{cc} 2 & 1 \\ 1 & 4 \end{array}\right)$

Table 8. Mixing Proportions, Mean vectors and Covariance matrices for simulation studies

The table depicts the four sets of means and covariance matrices where the first two sets have unequal covariance matrices while the other have equal covariance matrices.

are carried out and are reported in table (13). Results indicate significant outcomes meaning to the difference between the AUC's of MROC and mMROC models. This relates to that the AUC obtained through a mixture MROC with three components provides better information about the classifier by having more accuracy than the two component MROC model. Under simulations, the sets A & C will speak about the better classification scenario and sets B & D relates to worst class scenario. The AUC comparisons were found to be significant with sets A & C as they have minimum overlapping areas. Further, due to the maximum overlapping areas exhibited by the sets B & D resulted insignificant p-values where the AUC's are close to each other.

The MROC and mMROC curves at each parameter combination for varying sample sizes are shown in figures (9) and (10).

Summary

In this work, the main aim is to address on the issue of constructing an ROC model when there exists multi-model patterns in the known class labels. In medical scenario, most of data sets exhibit such patterns. In such situations, before we proceed for ROC modelling, it is suggested to look for such multi-model patterns that might exist in the data, if any. An illustration of this kind is demonstrated and to model such patterns, mUROC and mMROC models The practical applicability are proposed. and necessary illustrations of these models are given using OGTT, Disk Hernia data sets and simulation studies. Based on the results, it is understood that proposed ROC forms notifies better accuracy when compared with Bi-Normal and MROC models. Basically, the Bi-Normal ROC model and Multivariate ROC models were developed for a two -class classifications, so when we have a multi-class situation, one cannot rely on these models. So, in such cases, the proposed mUROC and mMROC models will be of great help and the graphical depiction of the ROC curves can be presented in a 2-dimensional space rather than with a complex higher dimension. On the whole, it is always suggested that before proceeding to modelling ROC curves, it is good to take a look at the density patterns of the study variable(s), which in turn will help in explaining the true information between the classes and also provides good amount of "true" accuracy about the marker(s).

	Sample							MROC	parameters					
Set	size		$\hat{\mu}_{_{\rm H}}$			$\hat{\mu}_{D}$			$\hat{\Sigma}_{\mathrm{H}}$			$\hat{\Sigma}$	D	
	25	(10.10454 6.69076)	(17.66999 10.98911)	(2.00686 1.01319 1.01319 2.99530)	(21.10063 6.07393	6.07393 6.46281)
	50	(10.09904 6.70402)	(17.49570 10.96989)	(1.998400.995450.995453.01684)	(20.82954 6.02989	6.02989 6.34245)
А	100	(10.09583 6.70182)	(17.50580 10.95309)	(1.992810.993410.993412.98802)	(20.62403 5.97472	5.97472 6.31814)
	250	(10.09869 6.70271)	(17.49997 10.95445)	(1.98469 0.98826 0.98826 2.98348)	(20.55593 5.95532	5.95532 6.32159)
	500	(10.10112 6.70264)	(17.50082 10.95650)	(1.997490.999030.999032.99258)	(20.51221 5.93948	5.93948 6.31243)
	25	(10.09200 6.69816)	(10.10787 6.68796)	(1.985081.007121.007123.02619)	(1.99265 0.97685	0.97685 5.10118)
	50	(10.10507 6.69237)	(10.09890 6.70950)	(2.00315 1.00625 1.00625 2.98878)))	(1.98671 0.99043	0.99043 5.00878)
В	100	(10.10062 6.69737)	(10.09922 6.69334)	(2.00079 1.00461 1.00461 3.02001)	(2.00555 1.00879	1.00879 5.01907)
	250	(10.09578 6.69398)	(10.09922 6.69441)	(2.003100.998000.998002.99502)	(1.99577 0.99195	0.99195 4.99197)
	500	(10.10159 6.69711)	(10.10162 6.69796)	(2.00576 0.99948 0.99948 2.99706)	(1.99757 0.99724	0.99724 4.98736)
	25	(10.09550 6.69838)	(17.65491 10.98152)	(1.997511.002361.002363.99925)	(21.47138 6.15339	6.15339 5.35069)
	50	(10.10188 6.69930)	(17.49161 10.93902)	(1.986040.975910.975914.00349)	(20.95016 6.05907	6.05907 5.33796)
С	100	(10.09579 6.69797)))	(17.49404 10.95740)	(1.998420.982880.982884.00100)	(20.68148 5.98281	5.98281 5.31919)
	250	(10.09810 6.70231)	(17.49812 10.94947)	(2.00667 1.01494 1.01494 3.99923)	(20.55439 5.95898	5.95898 5.31075)
	500	(10.10036 6.70202)	(17.50165 10.94792)	(1.996610.990250.990253.98883)	(20.52614 5.95700	5.95700 5.34308)
	25	(10.10083 6.71296)	(10.10859 6.72491)	(1.96491 1.01037 1.01037 3.99736)	(1.99364 1.00184	1.00184 3.99035)
	50	(10.11116 6.69497)	(10.10245 6.69943)	(2.00147 1.01289 1.01289 3.99149)	(2.00248 0.99474	0.99474 3.93676)
D	100	(10.10983 6.70429)	(10.09890 6.69710)	(1.99516 0.98727 0.98727 3.97426)	(2.01451 1.00626	1.00626 4.00825)
	250	(10.09920 6.69115)	(10.09907 6.70065)	(1.997131.007141.007144.01377)	(1.99593 0.99290	0.99290 3.98825)
	500	(10.09996 6.69683)	(10.10075 6.69958)	(2.00361 0.99686 0.99686 3.99104)	(1.99583 0.99669	0.99669 3.98328)

Table 9. Estimated means and covariance matrices of MROC model at various sample sizes

The table depicts the estimates of Mean vectors and Covariance matrices of the four sets considered for simulation studies at various

samples sizes.

Table 10. Estimated mixing proportions, means and covariance matrices of proposed mMROC model at various sample sizes

G (Sample					mMR	OC parameter	°s				
Set	size	$\hat{\lambda}_1$	$\hat{\lambda}_2$	$\hat{\mu}_{_{\rm H}}$	$\hat{\mu}_{D_1}$	$\hat{\mu}_{D_2}$	$\hat{\Sigma}_{\mathrm{F}}$	ł	$\hat{\Sigma}$	D	$\hat{\Sigma}_{]}$	D ₂
	25	0.49127	0.50873	(10.10454 6.69076)	(13.21104 9.80754)	$\left(\begin{array}{c} 21.78593 \\ 12.07980 \end{array} \right)$	(2.00686 1.01319	1.01319 2.99530	$\left(\begin{array}{c} 1.98524 \\ 0.98058 \end{array}\right)$	0.98058 3.97023	$(\begin{array}{c} 1.96485\\ 1.02805\end{array})$	1.02805 6.18519
	50	0.49927	0.50073	(10.09904 6.70402)	$\left(\begin{array}{c}13.20086\\9.82320\end{array}\right)$	$\left(\begin{array}{c} 21.79054\\ 12.11657\end{array}\right)$	$\left(\begin{array}{c} 1.99840\\ 0.99545 \end{array} \right)$	0.99545 3.01684)	$\left(\begin{array}{c} 1.99731 \\ 0.99534 \end{array}\right)$	0.99534 3.96877)	(2.01656 1.01002	1.01002 6.02197)
А	100	0.49929	0.50071	(10.09583 6.70182)	$\left(\begin{array}{c}13.21230\\9.80711\end{array}\right)$	$\left(\begin{array}{c} 21.79931\\ 12.09907 \end{array}\right)$	(1.99281 0.99341	0.99341 2.98802	$\left(\begin{array}{c} 1.99454\\ 0.99022 \end{array} \right)$	0.99022 3.98587	(2.01246 1.02165	1.02165 5.99822)
	250	0.49965	0.50035	(10.09869 6.70271)	$\left(\begin{array}{c} 13.20145\\ 9.80519\end{array}\right)$	$\left(\begin{array}{c} 21.79850\\ 12.10371 \end{array}\right)$	$\left(\begin{array}{c} 1.98469\\ 0.98826 \end{array} \right)$	0.98826 2.98348)	(2.00966 0.99165	0.99165 3.99214)	$\left(\begin{array}{c} 1.99881\\ 0.99844 \end{array} \right)$	0.99844 6.00218)
	500	0.49998	0.50002	(10.10112 6.70264)	$\left(\begin{array}{c} 13.20264\\ 9.80916 \end{array}\right)$	$\left(\begin{array}{c} 21.79900\\ 12.10383 \end{array} \right)$	$\left(\begin{array}{c} 1.99749\\ 0.99903 \end{array} \right)$	0.99903 2.99258	$\left(\begin{array}{c} 1.99987\\ 0.98793\end{array}\right)$	0.98793 3.99468	(2.00201 1.00886	1.00886 5.99283
	25	0.50093	0.49907	(10.09200 6.69816)	$\left(\begin{array}{c}10.11433\\6.70119\end{array}\right)$	$\left(\begin{array}{c}10.10190\\6.67575\end{array}\right)$	(1.98508 1.00712	1.00712 3.02619	(2.00313 1.00074	1.00074 3.99882)	$\left(\begin{array}{c} 1.97246\\ 0.95437 \end{array} \right)$	0.95437 6.11482)
	50	0.50105	0.49895	(10.10507 6.69237)	$\left(\begin{array}{c}10.10010\\6.72143\end{array}\right)$	$\left(\begin{array}{c}10.09769\\6.69758\end{array}\right)$	(2.00315 1.00625	1.00625 2.98878)	$\left(\begin{array}{c} 1.98096\\ 0.99777 \end{array} \right)$	0.99777 4.05185	$\left(\begin{array}{c} 1.99221\\ 0.98156 \end{array} \right)$	0.98156 5.96817)
В	100	0.50843	0.49157	(10.10062 6.69737)	$\left(\begin{array}{c}10.09895\\6.69230\end{array}\right)$	$\left(\begin{array}{c}10.09949\\6.69438\end{array}\right)$	(2.00079 1.00461	1.00461 3.02001	$\left(\begin{array}{c} 1.99755\\ 1.00190 \end{array} \right)$	1.00190 4.01536)	(2.01510 1.01556	1.01556 6.02106
	250	0.50501	0.49499	(10.09578 6.69398)	$\left(\begin{array}{c}10.09937\\6.69374\end{array}\right)$	$\left(\begin{array}{c}10.09906\\6.69509\end{array}\right)$	(2.00310 0.99800	0.99800 2.99502	$\left(\begin{array}{c} 2.00751\\ 0.99489 \end{array} \right)$	0.99489 3.98802	$\left(\begin{array}{c} 1.98479\\ 0.98985 \end{array} \right)$	0.98985 5.99721)
	500	0.50860	0.49140	(10.10159 6.69711)	$\left(\begin{array}{c}10.10154\\6.69905\end{array}\right)$	$\left(\begin{array}{c} 10.10169\\ 6.69688 \end{array}\right)$	(2.00576 0.99948	0.99948 2.99706)	$(\begin{array}{c} 1.99830\\ 1.00449 \end{array})$	1.00449 3.98450)	$(\begin{array}{c} 1.99688\\ 0.99017\end{array})$	0.99017 5.98912)
	25	0.49092	0.50908	(10.09550 6.69838)	$\left(\begin{array}{c} 13.15651\\ 9.79545\end{array}\right)$	$\left(\begin{array}{c} 21.80729\\ 12.07636\end{array}\right)$	(1.99751 1.00236	1.00236 3.99925	$\left(\begin{array}{c} 1.97665\\ 0.96691 \end{array} \right)$	0.96691 3.97046)	(2.04887 1.07399	1.07399 4.03201)
	50	0.49802	0.50198	(10.10188 6.69930)	$\left(\begin{array}{c}13.18281\\9.78740\end{array}\right)$	$\left(\begin{array}{c} 21.80040\\ 12.09065 \end{array}\right)$	$\left(\begin{array}{c} 1.98604\\ 0.97591 \end{array} \right)$	0.97591 4.00349	(2.00042 1.00927	1.00927 4.05739	(2.00497 0.97327	0.97327 3.90720)
С	100	0.49907	0.50093	(10.09579 6.69797)	$\left(\begin{array}{c}13.19387\\9.80767\end{array}\right)$	$\left(\begin{array}{c} 21.79421\\ 12.10713 \end{array}\right)$	$\left(\begin{array}{c} 1.99842\\ 0.98288 \end{array} \right)$	0.98288 4.00100)	(2.00003 0.98627	0.98627 3.97636	$\left(\begin{array}{c} 2.00741 \\ 0.98873 \end{array} \right)$	0.98873 3.98610
	250	0.49923	0.50077	(10.09810 6.70231)	$\begin{pmatrix} 13.19969 \\ 9.79946 \end{pmatrix}$	$\left(\begin{array}{c} 21.79656\\ 12.09948 \end{array}\right)$	(2.00667 1.01494	1.01494 3.99923	$\left(\begin{array}{c} 2.01018\\ 0.99708 \end{array} \right)$	$\left(\begin{array}{c} 0.99708\\ 3.98460 \end{array} \right)$	$\left(\begin{array}{c} 1.99708 \\ 0.99388 \end{array} \right)$	0.99388 3.98274)
	500	0.50112	0.49888	(10.10036 6.70202)	$\left(\begin{array}{c}13.20202\\9.79707\end{array}\right)$	$\left(\begin{array}{c} 21.80128\\ 12.09877 \end{array} \right)$	(1.99661 0.99025	0.99025 3.98883	(2.00753 0.99785	0.99785 4.01736	$(\begin{array}{c} 1.99722\\ 1.00055\end{array})$	1.00055 4.01531)
	25	0.50114	0.49886	(10.10083) 6.71296	$\left(\begin{array}{c}10.09875\\6.73157\end{array}\right)$	$\left(\begin{array}{c}10.11768\\6.71876\end{array}\right)$	(1.96491 1.01037	1.01037 3.99736	(2.00809 1.01407	1.01407 3.98593	$\left(\begin{array}{c} 1.98168\\ 0.98925 \end{array} \right)$	0.98925 3.98779
	50	0.50121	0.49879	(10.11116 6.69497)	$\left(\begin{array}{c}10.11027\\6.70019\end{array}\right)$	$\left(\begin{array}{c}10.09463\\6.69867\end{array}\right)$	(2.00147 1.01289&	1.01289 3.99149	$(2.01250 \\ 1.00685)$	1.00685 3.92122)	$\left(\begin{array}{c} 1.99582\\ 0.98459 \end{array} \right)$	0.98459 3.95430
D	100	0.50130	0.49870	(10.10983) 6.70429	$\left(\begin{array}{c} 10.09784\\ 6.70444\end{array}\right)$	$\left(\begin{array}{c} 10.09996\\ 6.68976 \end{array}\right)$	$\left(\begin{array}{c} 1.99516\\ 0.98727 \end{array} \right)$	0.98727 3.97426)	(2.01704 0.99471	0.99471 3.99336)	(2.01485 1.02091	1.02091 4.02971)
	250	0.50239	0.49761	(10.09920 6.69115)	$\left(\begin{array}{c}10.10184\\6.69570\end{array}\right)$	$\left(\begin{array}{c} 10.09630\\ 6.70561 \end{array}\right)$	(1.99713 1.00714	1.00714 4.01377)	$\left(\begin{array}{c} 1.98638\\ 0.98227 \end{array} \right)$	0.98227 3.99369	(2.00607 1.00420	1.00420 3.98474)
	500	0.50233	0.49768	(10.09996 6.69683)	$\left(\begin{array}{c}10.10272\\6.70330\end{array}\right)$	$\left(\begin{array}{c}10.09878\\6.69587\end{array}\right)$	(2.00361 0.99686	0.99686 3.99104	$\left(\begin{array}{c} 1.99838\\ 0.99824\end{array}\right)$	0.99824 3.97943)	(1.99366 0.99477	0.99477 3.98684)

The table depicts the estimates of mixing proportions, Mean vectors and Covariance matrices of the four sets considered for simulation

studies at various samples sizes.

Table 11. The linear combination	ations, cutoff points and measur	res of MROC for various sample sizes at

G .	Sample	MROC measures								
Set	size	U	с	FPR	TPR	J				
	25	$0.49503^*x_1^{} + 0.59497^*x_2^{}$	10.92129	0.09030	0.90970	0.81940				
	50	$0.48277^*x_1 + 0.57865^*x_2$	10.61223	0.09459	0.90541	0.81081				
А	100	$0.48848^*x_1^{} + 0.56082^*x_2^{}$	10.52059	0.09454	0.90546	0.81092				
	250	$0.48841^*x_1 + 0.55482^*x_2$	10.46879	0.09482	0.90518	0.81036				
	500	$0.48824^{*}x_{1}^{}+0.55245^{*}x_{2}^{}$	10.45343	0.09503	0.90497	0.80994				
	25	$0.01161^*x_1 - 0.00363^*x_2$	0.09753	0.42827	0.57173	0.14346				
	50	$\textbf{-0.0074}^* x_1 + 0.007456^* x_2$	-0.02238	0.44859	0.55141	0.10281				
В	100	$\textbf{-0.00012}^* x_1^{} \textbf{-0.00076}^* x_2^{}$	-0.00498	0.46483	0.53517	0.07034				
	250	$0.00184^*x_1 \text{ -} 0.00013^*x_2$	0.01831	0.47736	0.52264	0.04527				
	500	$-0.00007^*x_1 + 0.00015^*x_2$	-0.00402	0.48397	0.51603	0.03205				
	25	$0.47911^*x_1^{} + 0.60050^*x_2^{}$	10.88749	0.09416	0.90584	0.81169				
	50	$0.480569^* x_1^{} + 0.5695^* x_2^{}$	10.61016	0.09813	0.90187	0.80374				
С	100	$0.48418^* x_1^{} + 0.56287^* x_2^{}$	10.60277	0.09764	0.90236	0.80472				
	250	$0.48759^*x_1^{} + 0.55150^*x_2^{}$	10.55049	0.09807	0.90193	0.80386				
	500	$0.48937^*x_1^{} + 0.54786^*x_2^{}$	10.53425	0.09786	0.90214	0.80427				
	25	$0.00609^*x_1^{} + 0.00190^*x_2^{}$	0.10743	0.42781	0.57219	0.14438				
	50	$-0.00535^* x_1^{} + 0.00227^* x_2^{}$	-0.03975	0.45048	0.54952	0.09904				
D	100	$\textbf{-0.00498}^* x_1 \textbf{-0.00032}^* x_2$	-0.05225	0.46457	0.53543	0.07085				
	250	-0.00113 $^{*}x_{1}$ +0.00254 $^{*}x_{2}$	0.00521	0.47755	0.52245	0.04490				
	500	$0.00008^* x_1 + 0.00065^* x_2$	0.00520	0.48425	0.51575	0.03150				

The table contains the best linear combinations and AUC measures of the MROC model for simulation studies at various samples sizes.

C -4	C1	mMROC measures										
Set	Sample size	U ₁	U ₂	τ_1	τ2	FPR	TPR	J _M				
	25	$1.49517^*x_1 + 0.59813^*x_2$	$5.23432^{*}x_{1}^{-}-0.53711^{*}x_{2}^{-}$	22.54886	85.80980	0.05549	0.94451	0.88901				
	50	$1.36389^*x_1 + 0.56187^*x_2$	$4.77574^*x_1^{} - 0.48233^*x_2^{}$	20.57286	78.23666	0.05726	0.94274	0.88548				
А	100	$1.34323^*x_1 + 0.53634^*x_2$	$4.64862^*x_1^- \textbf{-} 0.46568^*x_2^-$	20.08756	76.18369	0.05770	0.94230	0.88460				
	250	$1.30705^*x_1 + 0.52759^*x_2$	$4.55944^*x_1^ 0.44248^*x_2^-$	19.55593	74.92880	0.05866	0.94134	0.88269				
	500	$1.29827^*x_1 + 0.52402^*x_2$	$4.54416^* x_1^{} - 0.44578^* x_2^{}$	19.42758	74.60272	0.05892	0.94108	0.88216				
	25	$0.01557^*x_1 + 0.00120^*x_2$	$-0.00148^* x_1 - 0.00398^* x_2$	0.22413	-0.03913	0.40087	0.59913	0.19826				
	50	-0.01024* x_1 +0.01111* x_2	$0.00141^* x_1 - 0.00310^* x_2$	-0.00584	-0.00044	0.43259	0.56741	0.13483				
В	100	$0.00121^* x_1 - 0.00179^* x_2$	$-0.00124^*x_1 + 0.00090^*x_2$	0.01106	-0.00308	0.45294	0.54706	0.09412				
	250	$0.00202^* x_1 - 0.00018^* x_2$	$0.00102^*x_1 + 0.00013^*x_2$	0.02381	0.00169	0.47001	0.52999	0.05998				
	500	-0.00018 $^{*}x_{1}$ +0.00063 $^{*}x_{2}$	$0.00036^*x_1 - 0.00064^*x_2$	-0.00526	0.00611	0.47868	0.52132	0.04265				
	25	$1.51072^*x_1 + 0.51681^*x_2$	$5.27942^*x_1 - 0.71801^*x_2$	22.02856	84.23608	0.05761	0.94239	0.88479				
	50	$1.40755^*x_1 + 0.46361^*x_2$	$4.87680^* x_1^{} \text{-} 0.60704^* x_2^{}$	20.27628	78.66009	0.05941	0.94059	0.88117				
С	100	$1.36271^*x_1 + 0.46260^*x_2$	$4.70298^*x_1 - 0.57506^*x_2$	19.71705	75.95861	0.05947	0.94053	0.88105				
	250	$1.33538^*x_1 + 0.44649^*x_2$	$4.62757^*x_1 - 0.57342^*x_2$	19.25432	74.73332	0.06041	0.93959	0.87919				
	500	$1.33518^*x_1 + 0.44507^*x_2$	$4.60400^* x_1 - 0.56903^* x_2$	19.23025	74.37660	0.06033	0.93967	0.87934				
	25	$-0.00264^{*}x_{1} + 0.00735^{*}x_{2}$	$0.01219^{*}x_{1}^{}$ - $0.01394^{*}x_{2}^{}$	0.23539	-0.21874	0.40334	0.59666	0.19332				
	50	$0.00039^*x_1 + 0.00192^*x_2$	-0.00699*x ₁ -0.00046*x ₂	0.03735	-0.07294	0.43399	0.56601	0.13202				
D	100	-0.00653* x_1 +0.00137* x_2	$0.00322^* x_1 - 0.00331^* x_2$	-0.04720	0.01146	0.45404	0.54596	0.09192				
	250	$0.00112^*x_1 + 0.00082^*x_2$	-0.00423 $^{*}x_{1}$ +0.00363 $^{*}x_{2}$	0.02053	-0.01882	0.47070	0.52930	0.05860				
	500	$0.00069^*x_1 + 0.00152^*x_2$	-0.00151*x ₁ -0.00157*x ₂	0.01925	-0.02596	0.47928	0.52072	0.04144				

Table 12. The linear combinations, cutoff points and measures of mMROC for various sample sizes at

The table contains the best linear combinations and AUC measures of the mMROC model for simulation studies at various samples sizes.

Set	Sample size	ÂUC	mÛÂÛC	Z _{AUC}	S _{ig.}
	25	0.95945	0.97649	1.66738	0.04878*
	50	0.95752	0.97684	2.05329	0.02002*
А	100	0.95779	0.97710	2.86499	0.00209*
	250	0.95776	0.97682	4.45766	0.00000^{*}
	500	0.95778	0.97676	6.41528	0.00000^{*}
	25	0.59953	0.63591	0.04851	0.48065 ^{NS}
	50	0.57187	0.59391	0.03286	0.48689 ^{NS}
В	100	0.54937	0.56600	0.26207	0.39663 ^{NS}
	250	0.53184	0.54222	0.01249	0.4950 ^{NS}
	500	0.52256	0.53005	0.01203	0.49520 ^{NS}
	25	0.95859	0.97517	1.65011	0.04946*
	50	0.95667	0.97528	1.83489	0.03326*
С	100	0.95757	0.97587	1.92829	0.02691*
	250	0.95750	0.97557	2.61752	0.00443*
	500	0.95759	0.97576	3.66464	0.00012^{*}
	25	0.60044	0.63270	0.01908	0.49239 ^{NS}
	50	0.56952	0.59206	0.00551	0.49780^{NS}
D	100	0.54992	0.56457	0.00344	0.50137 ^{NS}
	250	0.53170	0.54133	0.00105	0.50042 ^{NS}
	500	0.52226	0.52926	0.00002	0.50001 ^{NS}

Table 13. The estimated AUC's of MROC and mMROC curve along with values at various sample sizes

*Significant;

NS, Not significant

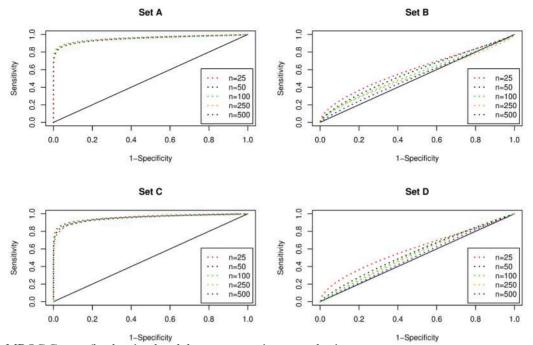
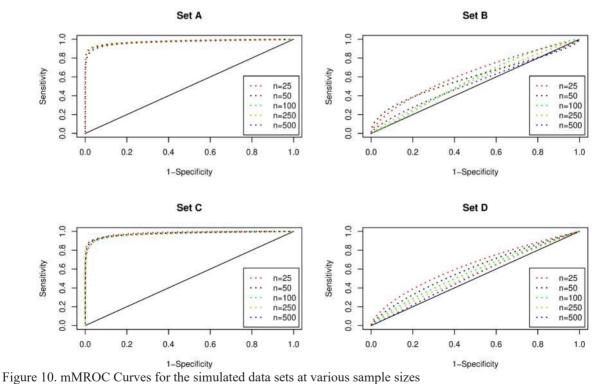


Figure 9. MROC Curves for the simulated data sets at various sample sizes The MROC curves for the four sets are shown here. The first and third sets are the examples of best cases while the second and fourth are the worst cases of classification.



The mMROC curves for the four sets are shown here. The first and third sets are the examples of best cases while the second and fourth are the worst cases of classification.

References

 Green DM, Swets JA. Signal detection theory and psychophysics. New York: Wiley. 1966.

 Egan JP. Signal Detection Theory and ROC Analysis "Academic Press. New York. 1975.

3. England WL. An exponential model used for optimal threshold selection on ROC Curves. Med Deci Mak, 1988;8:120-131.

4. Campbell G, Ratnaparkhi MV. An application of Lomax distributions in receiver operating characteristic (ROC) curve analysis. Comm in Stat -Theory and Methods. 1993;22:1681-1687.

5. Hussain E. The bi-gamma ROC curve in a straightforward manner. J Bas & App Sci. 2012;8:309-314.

6. Mossman D, Peng H. Using dual beta distributions to create "proper" ROC curves based on rating category data. Med Deci Mak. 2016;36:349-365.

7. Balaswamy S, Vardhan RV, Sarma, KVS. The Hybrid ROC Curve and its Divergence Measures for Binary Classification. Int J Stat in Med Res. 2015a;4(1):94–102.

8. Balaswamy S, Vardhan RV. Estimation of Confidence Intervals of a GHROC Curve in the presence of Scale and Shape parameters. Res J Math and Stat Sci. 2015b;3(10):4-11.

9. Balaswamy S, Vardhan RV. ROC curve Estimation using the combination

of Generalized Half Normal and Weibull distributions. J Indian Society for Prob and Stat. 2016a;17(1):11-23.

Balaswamy S, Vishnu Vardhan R.
 An Anthology of Parametric ROC Models.
 Research & Reviews: J Sta. 2016b;5(2):32-46.

11. Su JQ, Liu JS. Linear combinations of multiple diagnostic markers. J Ame Stat Assoc. 1993;88:1350-1355.

12. Schisterman EF, Faraggi D, Reiser B. Adjusting the generalized ROC curve for covariates. Stat in Med. 2004;23:3319-3331.

13. Yuan Z, Ghosh D. Combining multiple biomarker models in logistic regression. Biometrics. 2008;64:431-439.

14. Yin J, Tian L. Optimal linear combinations of multiple diagnostic biomarkers based on Youden index. Stat in Med. 2014;33:1426-1440.

15. Sameera G, Vardhan, RV, Sarma KVS. Binary classification using multivariate receiver operating characteristic curve for continuous data. J Biopha Stat. 2016;26:421-431.

16. Lasko TA, Bhagwat JG, Zou KH, Ohno-Machado L. The use of receiver operating characteristic curves in biomedical informatics. J Biomed Inf. 2005;38(5):404-415.

17. Dempster AP, Laird NM, Rubin DB. Maximum likelihood from incomplete data via the EM algorithm. J Royal Stat Soc: Ser B (Meth). 1977;39:1-22. 18. Titterington DM, Smith AFM, Makov UE. Statistical analysis of finite mixture distributions (Vol. 198). John Wiley & Sons Inc. 1985.

McLachlan G, Peel D. Finite MixtureModels. John Wiley and Sons, Inc., New York.2000.

20. Krzanowski WJ, Hand DJ. ROC curves for continuous data. Crc Press. 2009.

21. Anderson TW, Bahadur RR. Classification into two multivariate normal distributions with different covariance matrices. The Ann Math Sta. 1962;33:420-431.

22. Mossman D. Three-way rocs. Med Deci Mak. 1999;19:78-89.

23. Srinivasan A. Note on the location of optimal classifiers in n-dimensional ROC space. 1999.

24. Hand DJ, Till RJ. A simple generalisation of the area under the ROC curve for multiple class classification problems. Mach lear, 2001;45:171-186.

25. Ferri C, Hernandez-Orallo J, Salido MA. Volume under the ROC surface for multiclass problems. In Eur conf mach lear, Spri, Berlin, Heidelberg, 2003:108-120.

26. He X, Frey EC. The meaning and use of the volume under a three-class ROC surface (VUS). IEEE Trans Med Imag. 2008;27(5):577-588.

27. Kang L, Tian L. Estimation of the

volume under the ROC surface with three ordinal diagnostic categories. Comp Stat & Dat Ana. 2013;62:39-51.

28. Liu S, Zhu H, Yi K, Sun X, Xu W, Wang C. Fast and Unbiased Estimation of Volume Under Ordered Three-Class ROC Surface (VUS) With Continuous or Discrete Measurements. IEEE Access. 2020;8:136206-136222.

29. Youden WJ. Index for rating diagnostic tests. Cancer. 1950;3:32-35.

30. Perkins NJ, Schisterman EF. The Youden Index and the optimal cut-point corrected for measurement error. J Math Meth in Biosci. 2005;47:428-441.

31. Perkins NJ, Schisterman EF. The inconsistency of "optimal" cutpoints obtained using two criteria based on the receiver operating characteristic curve. Ame J Epid. 2006;163:670-675.

32. Leal12 J, Oliveira M, Sanches12 JM. Analysis of cut-off criteria in ROC curve for endarterectomy decision making. 2011.

33. da Rocha Neto AR, de Alencar Barreto G. On the application of ensembles of classifiers to the diagnosis of pathologies of the vertebral column: A comparative analysis. IEEE Lat Am Trans. 2009;7:487-496.